Discrete Choice & Applications

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Outline

Background

- Random utility models
- Definitions, equivalence properties
- Multinomial logit (MNL) models
- BTL, Plackett–Luce, Nested logit models
- k-RUMS, RUMS with ties
- Independence of irrelevant attributes

Algorithms

- Representations
- Compression and coresets
- Fitting
- Learning
- Special Cases of RUMs

Applications

- Connections to deep networks
- Network formation
- Document ranking
- User segmentation
- Recommendation systems
- Social choice and voting

Background

- Discrete choice
- Random utility models (RUMs)
- Definitions, equivalence properties
- Multinomial logit (MNL) models; BTL
- Independence of irrelevant attributes
- PCMC, Nested logit models, k-RUMS

Discrete choice



Where shall we eat tonight?

Discrete choice



Discrete choice



How do we get there?

Discrete choice: Factors



DistancePrice

• Cuisine type

Quality

- Time since last visit
- Companion opinion



Discrete choice: Repeat consumption



Most items we consume are not for the first time

- Sometimes go for reliability
- Sometimes go for novelty

Each day ...

Goal of discrete choice



Explain rational choice among discrete alternatives

Discrete choice as a field of study

- Important model in behavioral economics, social sciences, machine learning, etc
- Widely used in studying consumer demand in practice
- Especially important in online/interactive settings (search results, product alternatives, recommendations, etc)

• Daniel McFadden, 2000 Economics Nobel Prize

"for his development of theory and methods for analyzing discrete choice"

Modeling discrete choice

Universe = $[n] = \{1, ..., n\}$

Slates = non-empty subsets of [n]

Model. A function f: slate \rightarrow distribution over slate

- Captures uncertainty
- Can codify rational behavior

S and T highly overlap \Rightarrow f(S) and f(T) may be related



Example



Random utility model (RUM) [Marschak 1960]

- There is a distribution U on utility vectors $\{ [n] \rightarrow \mathbb{R} \}$
- Each user is drawn from U and will choose highest utility option in a slate

Utility vectors

Given a universe of n items, the user samples a utility vector $(u_1, ..., u_n)$ from a joint continuous distribution U



Given a slate S, the user will select the slate item with largest perceived utility



RUMs

- Continuous distribution U on utility vectors { [n] → ℝ }
 For simplicity, assume no ties
- Each user is (u₁, ..., u_n) ~ U iid and will choose highest utility option in a slate T (ie, argmax_{t∈T} u_t)
- Highly overlapping subsets will be related
 Eg, Pr[j | T] ≥ Pr[j | T ∪ {i}] for j ∈ T and i ∉ T
- Regularity: $Pr[j | T] \ge Pr[j | S]$, when $S \supseteq T$
- Rational behavior \Rightarrow order of utilities determines choice

Permutation process

- There is a distribution P on permutations $\{ [n] \leftrightarrow [n] \}$
- Each user is a permutation π ~ P and will choose highest ranked option, according to π, in a slate

Permutations

Given a universe of n items, the user samples a permutation $\pi = (u_{i1} > u_{i2} > \cdots > u_{in})$ from a distribution P



Equivalence

• Given a utility vector u, we can sort the items by utility, to obtain an equivalent permutation $\boldsymbol{\pi}$







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We can transform a RUM defined by a distribution U over utility vector into an equivalent RUM defined by a distribution over permutations P and vice versa

Permutations to winners



1

Discrete Choice

Winner distribution

Assume a universe [n] and a distribution on the permutations of [n]

Given a slate $S \subseteq [n]$, let $D_{S}(i)$ for $i \in S$ be the probability that a random permutation (ie, user) prefers i to every other element of S

Winner distribution (Eg)



Winner distribution (Eg)



Oracles for RUMs

Given a slate S

- \circ max-sample(S): picks an unknown random permutation π , and returns the element of S with maximum rank in π
- \circ max-dist(S): returns D_S(i), for all i ∈ S, ie, the probability that i wins in S given a random permutation

Oracles for RUMs (Eg)



- max-dist(S): return D_s = (0.2, 0.05, 0.75)

Multinomial logit (MNL) [Bradley & Terry 1952; Luce 1959]

Classical special case of RUMs

Model. Given a universe [n] of items and a positive weight a_j for each item $j \in [n]$

For a subset (slate) S of [n], the probability of choosing j in slate S is proportional to w_i

Pr[choosing j from S] = $w_j / \Sigma_{k \in S} w_k$

Permutations from an MNL (Eg)



Pick the next item in the permutation at random between the remaining ones, with probability proportional to its weight

MNLs = RUMs + specific noise

Assume each item j has an absolute "true quality" V_i

Model. Each user deviates from this by random noise ε_j and so the actual utility of user for item j is $u_j = V_j + \varepsilon_j$

 $Pr[user chooses j] = Pr[\forall k \neq j, V_j + \varepsilon_j > V_k + \varepsilon_k]$

Suppose ε_i 's are iid



A convenient choice of noise

Suppose $Pr[\varepsilon] = exp(-(\varepsilon + exp(-\varepsilon)))$

- Gumbel distribution
- Models the distribution of the maximum of samples (from various distributions)

Pr[user chooses j], by simple integration,

=
$$\Pr[j = \operatorname{argmax} \{V_k + \varepsilon_k\}] \propto \exp(V_j)$$

 \Rightarrow Pr[user chooses j from S] = exp(V_i) / $\Sigma_{k\in S}$ exp(V_k)

Multinomial regression gives identical choice probabilities to RUM with Gumbel-distributed noise!

Including features in MNL

We can make $V_{\mbox{\scriptsize j}}$ to depend on item features or user features or both

Suppose $V_j = \langle y_j, x \rangle$, where y_j is item feature for item j and x is the user feature Suppose $V_j = \langle y, x_j \rangle$, where y is feature of an item and x_j is user feature for item j

Multinomial logit Pr[user chooses j from S] = $exp(\langle y_j, x \rangle) / \Sigma_{k \in S} exp(\langle y_k, x \rangle)$ Choice MNL Pr[user chooses j from S] = $exp(\langle y, x_j \rangle) / \Sigma_{k \in S} exp(\langle y, x_k \rangle)$

MNLs in machine learning

MNLs, or softmax layers, are common in ML

- Multi-class problems
- Dual encoders
- Mixture of MNLs are sometimes used

 $Pr[output class j] = exp(\langle \beta_{i}, input \rangle) / \Sigma_{k} exp(\langle \beta_{k}, input \rangle)$

Limitations of MNL

Assume positive weight w_a for each item $a \in [n]$

Options: {a, b}: $Pr[a | a \text{ or } b] = w_a / (w_a + w_b)$

Options: {a, b, c}: $Pr[a | a \text{ or } b] = w_a / (w_a + w_b)$

Relative likelihood of a versus b does not depend on other alternatives: Choices are Independent of Irrelevant Alternatives (IIA), aka Luce's Axiom of Choice

 $MNL \equiv choice with IIA$

MNLs are insufficient to capture common settings

Luce's axiom of choice

Pr[a | a or b] does not change when c is added to slate



"Menu effect" or "decoy effect" in practice

Stationary rational choice might not follow IIA

User Type 1: 50%	5	100
User Type 2: 25%	100	1
User Type 3: 25%	75	1



Stationary rational choice might not follow IIA

User Type 1: 50%	5	100	15
User Type 2: 25%	100	1	75
User Type 3: 25%	75	1	100


Mixture of MNLs

Modeling distinct populations with simple MNL is the problem

Allowing a mixture of population, with a population-specific MNL, can solve the problem

- New items need not cannibalize equally from all other items
- Eg, a new bus route affects only bus riders

2-MNL mixture

Given a universe [n] of items and positive weights u_j and v_j for each item $j \in [n]$

For a slate S, the probability of choosing $j \in S$ equals

$$\gamma \cdot u_j / \Sigma_{k \in S} u_k + (1 - \gamma) \cdot v_j / \Sigma_{k \in S} v_k$$

Uniform mixture when $\gamma = 1/2$

MNL mixtures can approximate arbitrarily well any RUM [McFadden & Train 2000]

The story so far

RUM

- General approach to characterize choice
- Harder to interpret (and learn)

MNL

- Captures only RUMs with IIA
- Easy and fast to optimize
- Easy to interpret



A brief history of IIA

R Duncan Luce formulated "Axiom of Choice" (1959)

- Arrow (1951) proved the Impossibility Theorem showing that IIA was one of several mutually incompatible properties of a social choice function
- Bradley and Terry (1952) introduced a pairwise comparison choice model
 - Studied by Zermelo (1920s)
 - Often called the BTL model

Later, many authors, notably McFadden, completed the story extending BTL to MNL

What if IIA is violated?

Situation is much more complex....

Most powerful models are:

- Mathematically complex
- Computationally intractable
- Sophisticated external representations of dependence

Practitioners with non-IIA data typically use "Nested Logit"

Problems with IIA revisited

0.1



Nested logit

Modeling the decision as a tree is a nested, sequential, or hierarchical logit model. It looks like a sequence of multinomial logits. [McFadden 78]



Nested logit: Connections to RUM & MNL

Model. Nested Logit (NL) selects an item by traversing tree from root, applying MNL at each level

Casting NL as RUM:

- Utility of each item is a priori fixed
- Each user's utilities are perturbed
- Perturbation is drawn from specific joint distribution

Power of NL:

- Pros: Captures hierarchical cannibalization cleanly; generalizes MNL
- Cons: Choices must separate cleanly into nests

MNL in graphs

Model. Define a Markov chain given a graph where each node u has score $s_u > 0$

• Transition according to MNL choice

$$M_{uv} = S_v / \Sigma_{w \in \text{neighbors}(u)} S_w$$

MNL in graphs (Eg)

Transition probability proportional to the score of the node

• Eg,
$$M_{ac} = s_{c} / (s_{b} + s_{c} + s_{d})$$

Transition probabilities are context dependent

• Eg,
$$M_{ac} = 0.01$$
, $M_{ac} = 0.91$



Pairwise choice Markov chain [Ragain & Ugander 2016]

Given pairwise winning matrix M and slate S

- Construct M_s by restricting M to rows and columns of S
- Make M_s stochastic
- Compute stationary π_s of M_s to yield choice probability π_s (i) for item i
- Can represent BTL
- Not a RUM in general
 Can violate regularity
- Has other nice properties





10% There are only a few types of users

- Support of the permutation distribution is small
- Pragmatic
- Computationally helpful

Computational problems in RUMs (Eg)

Goal. Learn D_s , for all $S \subseteq [n]$

How to learn the probability distributions governing the choice in a generic slate?

Assume oracle access to RUMs

Assuming large slates is less realistic

Quickly learning the winning distributions of the slates is important for applications

... but there are exponentially many slates!

Computational problems in RUMs (Eg)

Goal. Given pairwise winning matrix, find the closest RUM

Head-to-head contests, online experiences comparing one item and an alternative

Algorithms

- Representations
- Compression and coresets
- Fitting
- Learning
- Special cases of RUMs

RUM Representations

- How to represent RUMs?
- How do different representations change the computational costs of various RUM tasks (e.g., fitting, learning)?

Given the full set of n items, the user samples a utility vector (u₁, u₂, ..., u_n) from a joint continuous distribution U



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Permutations

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This way, we can transform a RUM defined by a distribution *U* over utility vectors into an equivalent RUM defined by a distribution over permutations *P*, and vice versa.

Representations

- Is any of these two representations preferable for the tasks we are interested in, e.g.,
 - 1. storing/sketching a RUM,
 - 2. fitting a RUM, or
 - 3. learning a RUM?

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- But there are only finitely many permutations... thus, we can perfectly represent a RUM with *n*! many scalars

$$\pi_1 = n > n - 1 > ... > 2 > 1$$

 $\pi_2 = n > n - 1 > ... > 1 > 2$

$$\pi_{n!} = 1 > 2 > \dots > n - 1 > n$$

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- But there are only finitely many permutations... thus, we can perfectly represent a RUM with *n*! many scalars
- Can this representation be shrunk?
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$\{ \overbrace{\longleftarrow}, \quad \overleftarrow{\boxdot} \}, \{ \overbrace{\longleftarrow}, \quad \overleftarrow{\boxdot} \}, \dots, \{ \overbrace{\bigoplus}, \quad \overleftarrow{\bigtriangledown}, \quad \overleftarrow{\ominus} \}, \dots \}$

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Discrete Choice

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- Can RUMs be store more efficiently?



• Let *H* be a set of RUM items, and *s* an item not in *H*,



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- Let *H* be a set of RUM items, and s an item not in *H*,
- let $P_{H,s}$ be the probability that a random RUM permutation has the items of *H*, in any order, as its *|H|* top-most items, and it has s in position *|H|+1*.



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 $P_{\{4,6\},2} = \Pr_{\pi} [\pi \text{ begins with } 4 > 6 > 2 > \dots, \text{ or with } 6 > 4 > 2 > \dots]$



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- let $P_{H,s}$ be the probability that a random RUM permutation has the items of *H*, in any order, as its *|H|* top-most items, and it has s in position *|H|+1*.

The Head Distribution of item s is, then, $P_{\star,s}$, that is, the probability distribution over the **subset** of items that beat s in a random permutation (the **head** of s)

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$$D_S(s) = \Pr_{\pi} [s \text{ wins in } S \text{ with } \pi]$$

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= $\Pr_{\pi} [\pi \text{ begins with some set of elements disjoint} from S, and continues with s right after}]$
 $\pi = (x_{1} > \dots > x_{i} > s > \dots)$
with $\{x_{1}, \dots, x_{i}\} \cap S = \emptyset$

• The Head Distributions can *answer max-dist* queries:

$$D_{S}(s) = \Pr_{\pi} [s \text{ wins in } S \text{ with } \pi]$$

= $\Pr_{\pi} [\pi \text{ begins with some set of elements disjoint}$
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= $\sum_{T \subseteq [n] \setminus S} P_{T,s}$

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• As we said, $D_{[n]-H}(i) = \sum_{T \subseteq H} P_{T,i}$. Thus, $P_{H,i} = D_{[n]-H}(i) - \sum_{T \subseteq H} P_{T,i}$

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 - much smaller than that of permutations (which had n! dimensions), and
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 - much smaller than that of permutations (which had n! dimensions), and
 - having the same dimensionality of the input (the max-dist class $\{D_S(i)\}_{i\in S\subseteq [n]}$).
- While this is still very large, it cannot be improved if we want to *exactly* represent a RUM.

What is the Smallest Model for approximately representing a RUM?

Can we do lossy compression?

Approximate Representation



Discrete Choice

Approximate Representation



Discrete Choice

Let D be a RUM model on [n]Let D_s be the winner distribution of D on S

Model A ε -approximates D if, for each S \subseteq [n], $|D_S - A_S|_{TV} \leq \varepsilon$

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 $|D_S - A_S|_{TV}$ is the maximum gap between event probabilities in D_S and A_S

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 $|D_{S} - A_{S}|_{TV}$ is the maximum gap between event probabilities in D_{S} and A_{S}

Consider the event: "The winner in {1,2,3} is 1 or 3"

This event has probability 0.75 in D, and 0.77 in A

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Total Variation Distance

 $|D_{S} - A_{S}|_{TV}$ is the maximum gap between event probabilities in D_{S} and A_{S}

 $|(0.45, 0.25, 0.30) - (0.46, 0.23, 0.31)|_{TV} = (0.01 + 0.02 + 0.01) / 2 = 0.02$

Let D be a RUM model on [n]Let D_s be the winner distribution of D on S

Model A ε -approximates D if, for each S \subseteq [n], $|D_s - A_s|_{TV} \leq \varepsilon$

If we can find a model A,

 $\circ\,$ representable with few bits, and

• such that A ε -approximates D,

then we can efficiently sketch the RUM D to within Total Variation error $\pmb{\varepsilon}$

• [CKT21] proves that each RUM *D* on [*n*] can be sketched to within TV error ε , using $O(\varepsilon^{-2} n^2 \log n)$ bits.

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Let D' be the RUM obtained by imposing the uniform distribution on the sampled permutations



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THM: D' sketches *D* to an ε -TV error, w.p. 1-o(1)

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THM: D' sketches D to an ε-TV error, w.p. 1-o(1) THM: D' can be represented with

 $O(\varepsilon^{-2} n^2 \log n)$ bits

- [CKT21] proves that each RUM *D* on [*n*] can be sketched to within TV error ε , using $O(\varepsilon^{-2} n^2 \log n)$ bits.
- [CKT21] also proves that one cannot sketch the generic RUM
 D on [n] to within TV error 0.01, using o(n²) bits.

Size of Model vs Approximation Error



Storing a RUM

 RUMs are powerful choice models, whose perfect representations require an exponential number of bits,



Storing a RUM

- RUMs are powerful choice models, whose perfect representations require an exponential number of bits,
- but if one allows a tiny error, one can represent them efficiently with a number of bits bounded between $\Omega(n^2)$ and $O(n^2 \log n)$



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- In most practical applications, we do not observe the permutations, nor the utilities, of a RUM. We only observe the probability distributions over the winners of the slates.
- Recall that $D_S(i)$ is the probability that item *i* gets selected as the winner of slate *S*, for $i \in S \subseteq [n]$
- How to fit a RUM to these observed "winner distributions"?



- Let S_n be the set of permutations over $[n] = \{1, 2, ..., n\}$.
- Given a permutation $\pi \in S_n$, and a slate $S \subseteq [n]$, let $\pi(S)$ be the topmost item of S in π .

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If $\pi = 3 > 1 > 2$ and $S = \{1, 2\}$, then $\pi(S) = 1$

- Let S_n be the set of permutations over $[n] = \{1, 2, ..., n\}$.
- Given a permutation $\pi \in S_n$, and a slate $S \subseteq [n]$, let $\pi(S)$ be the topmost item of S in π .
- If there exists a RUM representing the winner distributions such a RUM can be directly obtained by solving the following LP:

$$\begin{cases} \sum_{\substack{\pi \in S_n \\ \pi(S)=i}} p_{\pi} = D_S(i) \quad \forall i \in S \subseteq [n] \\ \sum_{\substack{\pi \in S_n \\ \pi \in S_n}} p_{\pi} = 1 \\ p_{\pi} \ge 0 \qquad \forall \pi \in S_n \end{cases}$$

- This LP has n! many variables but it allows us to obtain a RUM compatible with the observed winner distributions in n^{O(n)} time.
- The existence of this LP (and of this finite fitting procedure) is another advantage of the combinatorial-based view of RUMs.

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In fact, by using a similar LP based on the Head distributions, one can obtain a "polytime" (2^{O(n)}) algorithm

One can also obtain the RUM "closest" to the input data, if no perfect RUM exists

- The "input" contains $\Omega(n 2^n)$ bits, thus a $n^{O(n)}$ algorithm (based on the permutation representation) is not too bad
- But, in many real-world situations, one does not have access to the winner distributions of all the slates but only to the winner distributions of slates of small size
- Can one obtain a polynomial-time fitting algorithm in that case?

• For simplicity, let us consider the case of slates of size 2.



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- The input to the fitting problem is then a matrix



 $D_{\{\widehat{\boldsymbol{m}}, \widehat{\boldsymbol{m}}\}}(\widehat{\boldsymbol{m}}) = 0.1$ $D_{\{\widehat{\boldsymbol{m}}, \widehat{\boldsymbol{m}}\}}(\widehat{\boldsymbol{m}}) = 0.6$

• • •

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- Many choice models have been proposed for representing tournament matrices:
 - Blade-Chest Chen & Joachims, WSDM '16
 - Majority Vote Makhijani & Ugander, WWW '19
 - Two-level model Veerathu & Rajkumar, NeurIPS '21

0 ...

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		Ä		⇒ ⇒ ⇒ □ p ₁			Ä	
		0.1	0.6	$\mathbf{P}_2 = \mathbf{P}_2$			0.1	0.6
Ä	0.9		0.3		Ä	0.9		0.3
	0.4	0.7		$ _{\Box} > \square > \square p_k$		0.4	0.7	
			Perfect Fit					

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		Ä		→ > → > → p ₁			Ä	
		0.12	0.6	\mathbf{P}_2			0.11	0.59
Ä	0.9		0.29		Ä	0.89		0.28
	0.4	0.7		$ \mathbf{p} > \mathbf{p} > \mathbf{p}_k$		0.41	0.72	
			Smallest "Frror" Fit					

• A Linear Program for minimizing the average TV-error:

$$\begin{cases} \min \sum_{\substack{1 \le i < j \le n \\ \sum m \in \mathbf{S}_n \\ \pi(\{i,j\}) = i \\ \epsilon_{i,j} \ge -e_{i,j} \\ \epsilon_{i,j} \ge -e_{i,j} \\ \epsilon_{i,j} \ge e_{i,j} \\ \sum m \in \mathbf{S}_n \\ p_{\pi} = 1 \\ p_{\pi} \ge 0 \end{cases} \quad \forall 1 \le i < j \le n \\ \forall 1 \le i < j \le n \\ \forall 1 \le i < j \le n \\ \forall 1 \le i < j \le n \\ \forall 1 \le i < j \le n \\ \forall 1 \le i < j \le n \\ \forall 1 \le i < j \le n \end{cases}$$

The LP has exponentially many variables!

- The Linear Program for minimizing the average TV-error has exponentially many variables, but only polynomially many constraints.
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Primal LP min c x under A x > bDual LP max by under $yA \le c$

Primal:

- 1 variable per permutation -
- 3 constraints per pair of items
- 1 extra constraint

Dual:

- 2 variables per pair of items 2 constraints per pair of items
 - 1 constraint per permutation
 - 3 variables per pair of items -
 - 1 extra variable

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O(n!) vars O(n²) constrs

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Strong Duality Theorem: $c x^* = b y^*$

- The Linear Program for minimizing the average TV-error has exponentially many variables, but only polynomially many constraints.
- Its dual then contains polynomially many variables.
- By means of the Ellipsoid method, if one could determine an unsatisfied dual constraint with a given solution, one would be able to optimize the primal and the dual - and, thus, find an optimal RUM.

- The Linear Program for minimizing the average TV-error has exponentially many variables, but only polynomially many constraints.
- Its dual then contains polynomially many variables.
- By means of the Ellipsoid method, if one could solve the dual Separation Oracle Problem, one would be able to optimize the primal and the dual - and, thus, find an optimal RUM.

Separation Oracle

- [ACKPT] observe that the separation oracle problem for the dual of the Pairwise RUM LP is equivalent to the Weighted Minimum Feedback Arc Set (WMinFAS) problem:
 - sort the vertices of a weighted directed graph, with weights bounded in [0,1], so that the total weight of the arcs directed left-to-right is minimized.
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D, C, A, B

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 - sort the vertices of a weighted directed graph, with weights bounded in [0,1], so that the total weight of the arcs directed left-to-right is minimized.
- MinFAS can be additively approximated to $O(\varepsilon n^2)$ in polynomial time for any constant $\varepsilon > 0$ [Frieze,Kannan,'99]

Approximate Separation Oracle

• [ACKPT] use this approximation algorithm for MinFAS to provide an Approximate Separation Oracle for the dual of the Pairwise LP. $\int \min_{1 \le i \le n} \epsilon_{i,j} \epsilon_{i,j}$

$$\operatorname{Primal} \left\{ \begin{array}{ll} \min \sum_{\substack{1 \le i < j \le n \\ \sum m \in \mathbf{S}_n \\ \pi(\{i,j\}) = i \\ \epsilon_{i,j} = i \\ \forall 1 \le i < j \le n \\ \forall 1 \le i < j \le n \\ \forall 1 \le i < j \le n \\ \forall 1 \le i < j \le n \\ \forall 1 \le i < j \le n \\ p_{\pi} = 1 \\ p_{\pi} = 0 \\ \forall \pi \in \mathbf{S}_n \end{array} \right.$$

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• [ACKPT] show that the Ellipsoid method, with this ASO, returns a RUM whose average TV-error is smaller than the min possible average TV-error plus ε , for any constant $\varepsilon > 0$.

Ellipsoid Method

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Ellipsoid Method

- The Ellipsoid method, while being a polynomial time algorithm, is inefficient in practice.
- [ACKPT] also show experimentally that the Approximate Separation Oracle can be used in practice, via a cutting-plane framework, for solving pairwise-RUM fitting on many instances.

Dataset	$\mid n$	avg. err.	lower bound on avg. err.
A5	16		
A9	12		
A17	13		0
A48	10		
A81	11		
SF	35	0.001408	0.001408
Jester	100	0.000461	0

- [CGKPT] show that the "pairwise" approach of [ACKPT] can be made to work on slates of size at most k = O(1):
 - $\circ\,$ to obtain this result, they study a more general LP, and give an algorithm for a generalized version of MinFAS

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Fitting so to minimize the **Average Error over the O(1)-slates**, can be ε -approximated in polynomial time

[ACKPT] show that the Approximate Separation Oracle for the **Maximum Error** over the 2-slates is NP-hard to approximate to within some additive constant

Learning a RUM

How well does a RUM fitted on slates of size at most *k* generalize to larger slates?

Streaming Services can test their users on small slates

$S = \{\textcircled{b}, \textcircled{b}, \textcircled{b}\}$



Streaming Services can test their users on small slates



Streaming Services can test their users on small slates



Streaming Services can test their users on small slates

$S = \{ \underbrace{b}_{0.2}, \underbrace{b}_{0.1}, \underbrace{b}_{0.7} \}$

Streaming Services can test their users on small slates

$S' = \{ b, b \}$



Discrete Choice

Streaming Services can test their users on small slates



Streaming Services can test their users on small slates

$S' = \{ \underbrace{b}_{0.4}, \underbrace{b}_{0.6} \}$

Streaming Services can test their users on small slates

$S'' = \{ \underbrace{b}_{0.8}, \underbrace{b}_{0.2} \}$

- Streaming Services can test their users on small slates
- It is impossible, though, to test the users on very large slates very few users would parse through a list of, say, 1000 movies to find their preferred one



 Streaming Services would love to pinpoint "gems" in their catalogues — items that are "most preferred" by a significant fraction of the user base



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Can they fit a RUM to what they observe on the small slates, and then use the RUM to guess the gems?

• In recent work, [CKGPT] show that – by accessing slates of size at most $O\left(\sqrt{n \cdot \ln \frac{1}{\epsilon}}\right)$ – one can approximate, to within an ε TV-error, the winner distribution of all slates of size at most *n*

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Accessing slates of size $O(\sqrt{n})$ exposes the structure of a RUM of *n* items to within a small error

In particular, accessing slates of size $O\left(\sqrt{n}\right)$ allows one to discover **gems**











• [CKGPT] also show that — if one can only access slates of size $o(\sqrt{n})$ — then one cannot guess if an item has probability at most ε , or at least 1- ε , in the slate {1,2, ..., n}.



This data is insufficient to guess whether there exists a **gem** in the catalogue
Generalization

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But, as we said, increasing the bound on the slate size to just above \sqrt{n} makes it possible to approximate all the winner distributions and, thus, to find **gems**

A fourth representation!

- This result shows that one can approximately represent a RUM with its winner distributions of slates of size at most $\approx \sqrt{n}$
- While the size of this representation is very large $(n^{O(\sqrt{n})})$, constructing the RUM this way gets us quite an improvement in the runtime $(2^{O(\sqrt{n} \ln n)} \text{ vs } 2^{O(n)})$ of RUM learning

RUM Representations

• RUM Representations

- Joint Utility Distribution
- (Light) Distribution over Permutations
- Head Distributions
- \circ Winner Distributions over slates of size at most $O(\sqrt{n})$
- They vary in their bit costs, and in the computational costs of various algorithmic tasks.
- Choose your representation wisely! :-)

Special Classes of RUMs

• Suppose that a RUM contains only *k* permutations in its support.

Discrete Choice

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 $P_{H,s}$ is the probability that a random permutation has the elements of *H*, in any order, as its |H| top-most elements, and that it has s in position |H|+1

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$$\mathsf{P}_{\{\mathbf{p}\},\mathbf{p}} > 0, \, \mathsf{P}_{\{\mathbf{p}\},\mathbf{p}} > 0$$

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$$P\{ \text{ }, \text{ } > 0, P\{ \text{ }, \text{ } > 0, P\{ \text{ }, \text{ } = \dots = P\{ \text{ }, \text{ } = 0 \}$$

- Suppose that a RUM contains only k permutations in its support.
- Then, for each cardinality *c*, there can be at most *k* pairs (*H*,*s*), with |H| = c, such that $P_{H,s}$ is non-zero.
- Thus, the formula $P_{H,s} = D_{[n]-H}(s) \sum_{T \subset H} P_{T,s}$

lets us learn the RUM with O(n k) max-dist queries.

Multinomial Logit MNL

- Classical special case of Random Utility Model
- Given a universe *U* of items and a positive weight *a_i* for each item *i* in *U*, the probability that *i* wins in the slate *S* is equal to

$$D_S(i) = \frac{a_i}{\sum_{j \in s} a_j}$$

For i = 1, ..., n-1, query the MNL using the slate $\{i, n\}$

obtaining the choice distribution

$$\left(rac{a_i}{a_i+a_n},rac{a_n}{a_i+a_n}
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$$\left(\frac{a_i}{a_i + a_n}, \frac{a_n}{a_i + a_n}\right)$$

System of
equations
$$\begin{array}{l}
\frac{a_n}{a_1 + a_n} = D_{1,n}(n) \\
\frac{a_n}{a_2 + a_n} = D_{2,n}(n) \\
\vdots \\
\frac{a_n}{a_{n-1} + a_n} = D_{n-1,n}(n) \\
\sum_{i=1}^n a_i = 1
\end{array}$$

For i = 1, ..., n-1, query the MNL using the slate $\{i, n\}$

Discrete Choice

Chierichetti, Kumar, Tomkins

For i = 1, ..., n-1, query the MNL using the slate $\{i, n\}$

obtaining the choice distribution

Querying *O(n)* slates of size 2, and solving this LP, gets us a valid set of weights

$$n \left(\frac{a_i}{a_i + a_n}, \frac{a_n}{a_i + a_n}\right)$$

$$a_n = D_{1,n}(n) \cdot (a_1 + a_n)$$

$$a_n = D_{2,n}(n) \cdot (a_2 + a_n)$$

$$\vdots$$

$$a_n = D_{n-1,n}(n) \cdot (a_{n-1} + a_n)$$

$$LP$$
Full Rank
$$LP$$

Mixture of MNLs

- MNL is insufficient to capture many practical settings
- 2-MNL mixture: Given a universe U of items and positive weights a_i and b_i for each item i in U

For a slate *S*, the probability of choosing *i* in *S* equals

$$\gamma \cdot \frac{a_i}{\sum_{j \in S} a_j} + (1 - \gamma) \cdot \frac{b_i}{\sum_{j \in S} b_j}$$

(Uniform mixture when $\gamma = 1/2$)

2-MNL Learning

- [CKT18] show that
 - Uniform 2-MNLs can be uniquely identified by the choice distributions of slates of sizes 2 and 3
 - There is a linear-time adaptive algorithm to learn the weights of uniform 2-MNLs using the choice distributions of slates of sizes 2 and 3

2-MNL Learning

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 - Uniform 2-MNLs can be uniquely identified by the choice distributions of slates of sizes 2 and 3
 - There is a linear-time adaptive algorithm to learn the weights of uniform 2-MNLs using the choice distributions of slates of sizes 2 and 3

Compare with general RUMs where, as we showed, one needs slates of size $O(\sqrt{n})$

2- and 3-Slates are sufficient

Theorem

For any uniform 2-MNL system, and for any set of 3 items $S = \{i, j, k\}$, the choice distributions of all the subsets of S determine uniquely the weights (up to rescaling) of *i*, *j*, *k* in each of the two MNLs.

Uniqueness

 This polynomial system induced by the choice distributions of the subsets of a generic set {*i*,*j*,*k*} has a unique solution

$$\begin{array}{l} \frac{a_{i}}{a_{i}+a_{j}} + \frac{b_{i}}{b_{i}+b_{j}} = 2D_{\{i,j\}}(i) \\ \frac{a_{i}}{a_{i}+a_{k}} + \frac{b_{i}}{b_{i}+b_{k}} = 2D_{\{i,k\}}(i) \\ \frac{a_{j}}{a_{j}+a_{k}} + \frac{b_{j}}{b_{j}+b_{k}} = 2D_{\{j,k\}}(j) \\ \frac{a_{i}}{a_{i}+a_{j}+a_{k}} + \frac{b_{i}}{b_{i}+b_{j}+b_{k}} = 2D_{\{i,j,k\}}(i) \\ \frac{a_{j}}{a_{i}+a_{j}+a_{k}} + \frac{b_{j}}{b_{i}+b_{j}+b_{k}} = 2D_{\{i,j,k\}}(j) \\ a_{i}+a_{j}+a_{k} = 1 \\ b_{i}+b_{j}+b_{k} = 1 \\ a_{i},a_{j},a_{k},b_{i},b_{j},b_{k} > 0 \end{array}$$

Discrete Choice

Algorithmic Implications

• Theorem

There exists an <u>adaptive</u> algorithm performing max-dist queries on <u>O(n) slates</u> of sizes 2 and 3, that reconstructs the weights of any uniform 2-MNL system on *n* elements.

Special Classes of RUMs

- RUMs supported on k permutations can be learned very efficiently
- Winner Distributions over slates of size at most $O(\sqrt{n})$ let you approximately represent any RUM
 - Winner Distributions over slates of size at most 2 let you represent any MNL
 - Winner Distributions over slates of size at most 3 let you represent any uniform 2-MNL
 - What about *k*-MNLs? Are slates of size *O(k)* sufficient for representation?

Applications

Applications

- ML Applications
- Geographic Choice
- Choice on Graphs
- Reconsumption

Conjoint Analysis

Initially developed by [Luce and Tukey 1964] – axiomatic formulation Picked up soon by marketers in late 60's, eg [Green and Rao 1971] B-T model:

Flavor	Mango	Chocolate
Price	2.95	3.95
Size	120g	200g

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Price	2.95	>	3.95
Size	120g		200g

Conjoint Analysis Outcomes



Widely used in marketing

"Like giving dynamite to babies"

Influential case study on Marriott Courtyard hotels

Relative performance

Courtyard by Marriott

ROOM PRICE PER NIGHT IS \$ 44.85

BUILDING SIZE, BAR/LOUNGE

Large (600 rooms) 12-story hotel with:

- Ouiet bar/lounge
- Enclosed central corridors and elevators
- All rooms have very large windows

LANDSCAPING/COURT

Building forms a spacious outdoor courtyard

- View from rooms of moderately landscaped courtyard with:
 - many trees and shrubs
 - the swimming pool plus a fountain
 - terraced areas for sunning, sitting, eating

FOOD

Small moderately priced lounge and restaurant for hotel guests/friends

- Limited breakfast with juices, fruit, Danish, cereal, bacon and eggs
- Lunch—soup and sandwiches only
- · Evening meal-salad, soup, sandwiches, six hot entrees including steak

HOTEL/MOTEL ROOM QUALITY

Quality of room furnishings, carpet, etc. is similar to:

- Hyatt Regencies
- Westin "Plaza" Hotels

Courtyard by Marriott

Attribute	Levels	Description	Part Worths	
Hotel Size	1	Small (125 rooms) 2-story hotel (.00)*	1.06	
	2	12-story (600 rooms) with large lobby,	0.00	
		meeting rooms, etc. (7.15)		
Corridor/View	1	Outside stairs and walkways to all	0.00	
		rooms. Restricted view. People		
		walking outside window. (.00)		
	2	Enclosed central corridors and stairs.	1.85	
		Unrestricted view. Rooms have	x	
		balcony or large window. (.65)	3	
Pool Location	1	Not in courtyard (.00)	0.00	
	2	In courtyard (.00)	1.37	
Pool Type	1	No pool (.00)	0.61	
	2	Rectangular pool (.45)	1.25	
	3	Freeform pool (.50)	0.29	
	4	Indoor/outdoor pool (.85)	0.00	
Landscaping	1	Minimal landscaping (.00)	0.81	
	2	Moderate landscaping (.10)	0.97	
	3	Elaborate landscaping (.50)	0.00	
Building Shape	1	"L" shape building with modest	0.00	
		landscaping (.00)		
	2	Building forms an outdoor landscaped	0.37	
		courtyard for sitting, eating,		
		sunning, etc. (.45)		

*Figure in parentheses after each description = price premium.

Softmax and discrete choice

A generic transformer (from [Vaswani et al 2017])



Softmax bottleneck

[Yang et al, 2018]

$$\begin{split} \mathbf{H}_{\theta} &= \begin{bmatrix} \mathbf{h}_{c_1}^{\top} \\ \mathbf{h}_{c_2}^{\top} \\ \vdots \\ \mathbf{h}_{c_N}^{\top} \end{bmatrix}; \ \mathbf{W}_{\theta} = \begin{bmatrix} \mathbf{w}_{x_1}^{\top} \\ \mathbf{w}_{x_2}^{\top} \\ \vdots \\ \mathbf{w}_{x_M}^{\top} \end{bmatrix}; \ \mathbf{A} = \begin{bmatrix} \log P^*(x_1|c_1), & \log P^*(x_2|c_1) & \cdots & \log P^*(x_M|c_1) \\ \log P^*(x_1|c_2), & \log P^*(x_2|c_2) & \cdots & \log P^*(x_M|c_2) \\ \vdots & \vdots & \ddots & \vdots \\ \log P^*(x_1|c_N), & \log P^*(x_2|c_N) & \cdots & \log P^*(x_M|c_N) \end{bmatrix} \\ F(\mathbf{A}) &= \{\mathbf{A} + \mathbf{A} \mathbf{J}_{N,M} | \mathbf{\Lambda} \text{ is diagonal and } \mathbf{\Lambda} \in \mathbb{R}^{N \times N} \}, \end{split}$$

Property 1. For any matrix \mathbf{A}' , $\mathbf{A}' \in F(\mathbf{A})$ if and only if $Softmax(\mathbf{A}') = P^*$. In other words, $F(\mathbf{A})$ defines the set of **all** possible logits that correspond to the true data distribution.

Property 2. For any $A_1 \neq A_2 \in F(A)$, $|rank(A_1) - rank(A_2)| \leq 1$. In other words, all matrices in F(A) have similar ranks, with the maximum rank difference being 1.

Goal of Languge Modeling: $\mathbf{H}_{\theta}\mathbf{W}_{\theta}^{\top} = \mathbf{A}'$.

Softmax bottleneck: rank of A' is at most the embedding dimension d

Softmax bottleneck – another view

Consider two nearby word representations – very difficult to separate

All "usage patterns" must be embedded into R^d

Mixture of softmax

$$P_{\theta}(x|c) = \sum_{k=1}^{K} \pi_{c,k} \frac{\exp \mathbf{h}_{c,k}^{\top} \mathbf{w}_{x}}{\sum_{x'} \exp \mathbf{h}_{c,k}^{\top} \mathbf{w}_{x'}}; \quad \text{s.t.} \quad \sum_{k=1}^{K} \pi_{c,k} = 1$$

MoS shows empirical wins over Softmax

The authors argue this is because it addresses the rank deficiency of the "softmax bottleneck"

Note that MoS is exactly mixed logit, and is there's equivalent to the full RUM family, where a user type is an assignment of "utilities" for each token

Utilities are a non-linear function of the context so far

Another take on the power of MNLs versus RUMs
Application: Geographic Choice

(or: where should we have dinner tonight?)

Where shall we eat tonight, revisited....



Discrete Choice

Chierichetti, Kumar, Tomkins

Some Factors in Restaurant Choice

Deciding where to go for dinner:

- Quality of the restaurant
- Distance from Hotel Michael
- Price
- Cuisine type
- \circ Ambience
- Time since last visit
- Opinions of dining companion(s)

Ο ...

Some data for this problem

Directions queries:

- Number of directions queries to US/Canadian restaurants in Google Maps
- Random sample of 15.5M queries to ~400K restaurants



Discrete Choice

Chierichetti, Kumar, Tomkins

Dataset

Directions queries:

- Number of directions queries to a US/Canadian restaurants in Google Maps
- Random sample of 15.5M queries to ~400K restaurants

Caveats:

- Not all visits have an associated directions search
 - Familiar locations
 - Spontaneous decisions
- Not all searches result in visits
 - Aspirational searches
 - Traffic & time estimates

Classical Discrete Choice Models

Recall our basic discrete choice model:

- Assign a score to each alternative
- Select with probability proportional to score

$$\Pr[x|A] = rac{w_x}{\sum_{y \in A} w_y}; w_x = e^{V_x}$$

Goal:

• Better understand the score

Score function

Today:

- Distance to the restaurant d
- Number of closer restaurants, rank: r
 - Captures density of restaurants
 - Acts as a proxy for the amount of competition
- Quality of particular restaurant: q
- Assume utility is linear in these features $V_x = d_x + r_x + q_x$

Not Today:

- Personal (user specific) preference
- Time since last visit
- Companions' desires

Imputed Rank Function



Imputed Distance Function



Results

Predict Likelihood on a held out test set:

Method	Likelihood
Uniform choice	1.1
Distance only model	3.9
Rank only model	4.6

Model

Fit both rank and distance functions by log-normals

• Four parameter model:

$$\mu_{\text{rank}}, \sigma_{\text{rank}}^{2}, \mu_{\text{distance}}, \sigma_{\text{distance}}^{2}$$

$$s_{i} = \frac{1}{r_{i}\sigma_{\text{rank}}} \exp\left(-\frac{(\ln r_{i} - \mu_{\text{rank}})^{2}}{2\sigma_{\text{rank}}^{2}}\right) \cdot \frac{1}{d_{i}\sigma_{\text{distance}}} \exp\left(-\frac{(\ln d_{i} - \mu_{\text{distance}})^{2}}{2\sigma_{\text{distance}}^{2}}\right)$$

Results

Predict Likelihood on a held out test set:

Method	Likelihood
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Results

Predict Likelihood on a held out test set:

Method	Likelihood
Uniform choice	1.1
Distance only model	3.9
Rank only model	4.6
Lognormal coefficient fit (4 parameters)	5.1
Non-parametric factored model	5.3

Quality Factor

- Quality is restaurant specific, makes the model much richer
- Learn it as the residual on ranks, distances
- Evaluation: correlation with critics' scores



Geographic Choice: what have we seen?

Multinomial Logistic Regression with buckets is a powerful technique to assess influence of features based on intensity

Captured interactions may give significantly different influence weights than feature correlations

Given the output of such models, it is possible to observe deeper structure

From this structure, we may find models that are far more parsimonious (why lognormal?)

These new models are much easier to fit when data is sparse

Application: Graphs

Ravi Kumar, Andrew Tomkins, Sergei Vassilvitskii and Erik Vee

[Ref: <u>WSDM 2015</u>]

Reverse Engineering a Markov Chain

Random Walks & Markov Chains

Markov Chains in Data Analysis:

- Simple, yet capture a lot of interactions
- Typically: compute & use the stationary distribution
- Beautiful theory with great applications

Examples:

- PageRank: Random surfer stationary distribution
- Translation: Use language models to build phrases

0 ...



mathworld.wolfram.com > ... > Markov Processes MathWorld A Markov chain is collection of random variables {X_t} (where the index t runs through 0, 1, ...) having the property that, given the present, the future is ...



Discrete Choice



Discrete Choice





Discrete Choice



Discrete Choice

Example:

- Items: videos
- Stationary Distribution: view counts

Why are some videos more popular:

- Better (higher quality) videos
- More frequently recommended

Today:

Disentangle these two reasons

Inverting a Markov Chain

Problem:

• Given a stationary distribution, find the Markov Chain that generated it.

Given:

- Graph G
- \circ Distribution π

Output:

• Transition Matrix *M* that generated it

Feasibility

Feasibility:

• Not always feasible



 π

Feasibility

Feasibility:

• Not always feasible



Definition:

- A directed graph is consistent if there is a flow that preserves the steady state.
- Any strongly connected graph with self loops is consistent

Theorem:

 $\circ~$ For any consistent graph, there exists a Markov chain with π as its stationary distribution.

Constraints

The problem is under-constrained:

- *n* constraints
- $\circ m n \gg n$ variables

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Approaches

• [Tomlin `03]: MaxEnt objective on variables (regularization)

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Approaches

- [Tomlin `03]: MaxEnt objective on variables (regularization)
- $\circ~$ [Today] Limit the degrees of freedom
- $\circ\;$ For each vertex v_i let s_i be its score. The Markov Chain is the function of the scores
- Scores express "quality" or "attractiveness"

From Scores to Transitions

Transition probability $M_{A \rightarrow C}$ depends on:

- \circ Score of the destination s_c
- \circ Parameter of the edge w_{AC}



Simplest Example

Weighted Random Walk:

- $\circ~$ All of the edge weights are set to 1
- Transition probability proportional to the score



$$M_{A \to C} = \frac{s_C}{s_B + s_C + s_D}$$

Simplest Example

Weighted Random Walk:

- $\circ~$ All of the edge weights are set to 1
- Transition probability proportional to the score



$$M_{A \to C} = \frac{s_C}{s_B + s_C + s_D}$$

• Transition probabilities are context dependent:

B
$$s_B = 100$$

A C $s_C = 10$
D $s_D = 1$
 $M_{A \to C} = 0.09$

Simplest Example

Weighted Random Walk:

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$$M_{A \to C} = \frac{s_C}{s_B + s_C + s_D}$$

• Transition probabilities are context dependent:

1

B
$$s_B = 100$$

 $M_{A \rightarrow C} = 0.09$
A C $s_C = 10$
 $M_{F \rightarrow C} = 0.91$
 F D $s_D = 1$
From Scores to Transitions

Transition probability $M_{A \rightarrow C}$ depends on:

- \circ Score of the destination s_c
- \circ Parameter of the edge w_{AC}
- \circ Call this function f

A D

Formally:

 $M_{A \to C} \propto f(s_C, w_{AC})$

 $M_{A \to C} = \frac{f(s_C, w_{AC})}{f(s_C, w_{AC}) + f(s_B, w_{AB}) + f(s_D, w_{AD})}$

From Scores to Transitions

Transition probability $M_{A \rightarrow C}$ depends on:

- \circ Score of the destination s_c
- \circ Parameter of the edge w_{AC}
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Formally:

 $M_{A \to C} \propto f(s_C, w_{AC})$

Sanity Check on : f

- Continuous in *S*
- $\circ~$ Monotone in ${\it S}$



From Scores to Transitions

Transition probability $M_{A \rightarrow C}$ depends on:

- \circ Score of the destination s_c
- \circ Parameter of the edge w_{AC}
- \circ Call this function f

Formally:

 $M_{A\to C} \propto f(s_C, w_{AC})$

Sanity Check on : f

- Continuous in *S*
- Monotone in *S*
- Unbounded in *S*

 $\lim_{s \to \infty} f(s, w) \to \infty$

$$\lim_{s_c \to \infty} M_{A \to C} = 1$$



Simplest Example

Weighted Random Walk:

- $\circ~$ All of the edge weights are set to 1
- Transition probability proportional to the score



$$M_{A \to C} = \frac{s_C}{s_B + s_C + s_D}$$

More Examples

Weighted Random Walk:

- All of the edge weights are set to 1
- Transition probability proportional to the score



$$M_{A \to C} = \frac{s_C}{s_B + s_C + s_D}$$

Seeking Similar Content:

 $\circ~$ Edge weight: similarity between two nodes $~M_{A \rightarrow C} \propto w_{AC} \cdot s_{C}$

More Examples

Weighted Random Walk:

- All of the edge weights are set to 1
- Transition probability proportional to the score



$$M_{A \to C} = \frac{s_C}{s_B + s_C + s_D}$$

Seeking Similar Content:

 $\circ~$ Edge weight: similarity between two nodes $M_{A
ightarrow C} \propto w_{AC} \cdot s_C$

Overall:

 Decide whether items are popular due to high scores (attract all of the incoming traffic) or due to location (attract a little bit from many locations)

Main Theorem

Given:

- A consistent input G, π
- \circ Monotone, continuous and unbounded function f

There exists:

- A unique set of scores s_1, \ldots, s_n
- $\circ~$ So that π is the stationary distribution induced by f
- Moreover, the scores can be found in polynomial time

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up to scaling

up to $(1 \pm \epsilon)$

 $\circ~$ Fix a set of scores ${\it S}$ and distribution π



- $\,\circ\,$ Fix a set of scores ${\it S}$ and distribution π
- \circ Let $q_i(s)$ be the expected mass at v_i starting with s using π



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- \circ Fix a set of scores s and distribution π
- \circ Let $q_i(s)$ be the expected mass at v_i starting with s using π
- Call a node underweight if $q_i(s) < (1 \epsilon)\pi_i$
- Algorithm:
 - Repeatedly increase scores of underweight nodes



- \circ Fix a set of scores s and distribution π
- \circ Let $q_i(s)$ be the expected mass at v_i starting with s using π
- Call a node underweight if $q_i(s) < (1 \epsilon)\pi_i$
- Algorithm:
 - Repeatedly increase scores of underweight nodes



- \circ Fix a set of scores *S* and steady state π
- \circ Let $q_i(s)$ be the expected mass at v_i starting with π using s
- Call a node underweight if $q_i(s) < (1 \epsilon)\pi_i$

Algorithm:

- \circ Start with $s_i^0 = 1/n$
- \circ For $t = 1, \ldots$
 - For each $v_i \in V$:
 - If v_i underweight: Set $s_i^t : q_i(s_{-i}^{t-1}, s_i^t) = (1 - \epsilon/2)\pi_i$
 - else:

$$\operatorname{Set} s_i^t = s_i^{t-1}$$

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 - else:

Set
$$s_i^t = s_i^{t-1}$$

Guaranteed to exist because f is monotone, continuous, unbounded & G is consistent

- Fix a set of scores s and steady state π
- Let $q_i(s)$ be the expected mass at v_i starting with π using **S**
- Call a node underweight if $q_i(s) < (1 \epsilon)\pi_i$



- Fix a set of scores s and steady state π
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Proof of Convergence

Key Lemma:

 \circ There is an explicit bound M such that $s_i^t \leq M$ for all i, t.

Proof of Convergence

Key Lemma:

• There is an explicit bound M such that $s_i^t \leq M$ for all i, t.

Proof Sketch:

- $\circ~$ Consider a set of scores that grows without bound
- $\circ~$ These scores all must be underweight (these are the only scores that increase)
- \circ Not all scores can be underweight (sum of underweight scores below 1)
- The scores growing without bound are taking all of the probability mass from those bounded
- $\circ~$ By consistency, this demand must be met, a contradiction.

Proof of Convergence

Key Lemma:

• There is an explicit bound M such that $s_i^t \leq M$ for all i, t.

Finishing the Proof:

 \circ Scores increase multiplicatively by factor of $(1+\epsilon/2)$

$$\circ M \text{ is bounded by } \left(\frac{n^2 W}{\epsilon p_{\min}}\right)^n$$
$$\circ \text{ Overall: } O\left(\frac{n^2}{\epsilon}\log\frac{nW}{\epsilon p_{\min}}\right) \text{ iterations suffice.}$$

But Does it Work...

Experimental Evaluation:

- Dataset: empirical transitions
- Input: Transition graph and the steady state distribution
- Output: Transition probabilities
- Metrics: LogLikelihood or RMSE

Datasets

Wiki:

- Navigation paths through wikipedia.
- About 200k transition pairs, 51k user traces over 4.6k nodes

Rest:

- Results of broad restaurant queries to Google.
- 100k transitions, 65k nodes

Entree:

- Chicago restaurant recommendation system from 90s
- 50k transitions, 27k nodes

Comedy:

- Given a pair of videos, predict which one is judged funnier
- 225k transitions, 75k nodes

Baselines

Popularity:

• Transition proportionally to the steady state distribution (score = pi)

Uniform:

Uniform over out-edges

Pagerank:

 $\circ~$ Transition proportionally to the node pagerank

Temperature:

• MaxEnt regularization approach

Inversion:

• Our algorithm

Results

RMSE Prediction:

	Popularity	Uniform	PageRank	Temp	Inversion
Wiki	1	0.65	0.83	0.65	0.57
Rest	1	1.17	1.39	1.21	0.59
Entree	1	0.69	1.01	0.56	0.42
Comedy	1	0.65	0.9	0.78	0.36

Application: Sequential Choice

Repeat consumption

Most of the items we consume are not for the first time

Sometimes go for reliability

Sometimes go for novelty

- Boredom
- New options

We focus on the repeated consumption, not the novel choice.



Repeat consumer choice

- Marketing studies
- Consumer behavior

Music listening experiment [Kahnx et al 97]

- Melioration/overconsumption: listen to favorite on each trial
- Maximization: preserve the high level of enjoyment

Possible explanations

- Difficulties in prediction of taste
- Users try to create the best memory (five flavors vs one flavor LifeSavers)
- Zen principles (pain vs pleasure)



Re-searching

Repeat queries in search logs [Teevan et al]

40% of queries are re-finding queries

Navigational queries are more likely to be repeated

• Information re-finding

Repeat behavior leads to easier prediction of which results will be clicked

Re-visiting web pages

Web page revisitation using browser logs [Adar et al]

50-80% of the web pages are revisited

Revisitation reasons

- Bookmarks/use as hub
- Track content change
- Backbutton

Types of revisitation

- Fast: shopping pages, references, traffic
- Medium: mail, forums, news, ...
- Slow: weekend activity, software updates, ...

Domains of reconsumption

Location checkins

- BrightKite
- Google+



Clicks

- Businesses on maps
- Restaurants on maps
- Wikipedia

Media

- Youtube
- Music videos
- Playlists from a radio station







Characterizing Reconsumption

Does it exist?

Distribution of the fraction of repeat consumption



Lifetime distributions

Do items have finite lifetimes?


Boredom

Do users get bored with repeat consumption?

- Marketers, advertisers care about this
- Churn/variety-seeking behavior





Summary of model



Discrete Choice

Three key factors

- How popular is the item?
- Time gap since it was last consumed
- How recently was it consumed?
- Can we develop a holistic mathematical framework powerful yet simple enough to explain patterns of reconsumption we observe in real data?

Recency model

Empirically, recency seems to play a strong role in reconsumption

Technical approach: Combine discrete choice model with "copying model" [Simon, 55] based on recency



Score-based model

Each item x has a score s_x

The score reflects the quality of the item

The score dictates the reconsumption pattern

Pick next item x using discrete choice, with probability:

$$\Pr[x|X] = rac{s_x}{\sum_{y \in A} s_y}$$

Combining Recency and Quality

Pr[d consumed next] ~



At position i, pick item x with probability:

$$rac{\sum_{j < i} I(x_j = x) w_{i-j} s_{x_j}}{\sum_{j < i} w_{i-j} s_{x_j}}$$

Stochastic gradient ascent

Alternating updates to scores and weights

Likelihoods (wrt hybrid model)

$egin{array}{l} s(\cdot) = \ w(\cdot) = \end{array}$	popularity	popularity	learned	uniform
	-	learned	uniform	learned
BRIGHTKITE GPLUS MAPCLICKS WIKICLICKS YOUTUBE	$\begin{array}{c} 0.375 \\ 0.587 \\ 0.383 \\ 0.503 \\ 0.636 \end{array}$	$\begin{array}{c} 0.617 \\ 0.801 \\ 0.931 \\ 0.724 \\ 0.677 \end{array}$	$\begin{array}{c} 0.637 \\ 0.794 \\ 0.414 \\ 0.687 \\ 0.924 \end{array}$	0.936 0.877 0.989 0.945 0.962

- Recency comes close to hybrid model
- Recency much better than quality
- Popularity seems to bring the models down even with recency

Combining Recency, Quality, and Time



At position i, pick item x with probability:

$$rac{\sum_{j < i} I(x_j = x) w_{i-j} s_{x_j} t_{t_i - t_j}}{\sum_{j < i} w_{i-j} s_{x_j} t_{t_i - t_j}}$$

Stochastic gradient ascent

Alternating updates to scores and weights

Learned time scores $T(t_i - t_i)$

- Learned time scores are complex
- Capture, e.g., cyclic behavior in check-in data.



Model Quality

ne a color	Learned scores				
Dataset	w	w and s	w and T		
BRIGHTKITE	0.91	0.92	0.98		
GPLUS	0.87	0.92	0.94		
LASTFM	0.99	0.99	1.00		
LASTFMARTISTS	0.96	0.96	1.00		
YOUTUBE	0.91	0.94	0.96		
YOUTUBEMUSIC	0.92	0.93	0.97		
MAPCLICKS	0.81	0.82	0.99		
WIKICLICKS	0.78	0.81	0.91		

- Score-only and Popularity-only not competitive
- Recency is most important feature
- Time is more important than item quality
- All model components bring some gain

Macroscopic observations

- 1. Eventual abandonment: item lifetime distributions are heavy-tailed and often finite.
- 2. Boredom: at the end of an item's life, gaps between consumptions increase monotonically.

Item lifetimes

Count lifetime: number of times an item is consumed.



Item lifetimes

Index lifetime: total number of items consumed between first and last consumption of a given item.



Item lifetimes

Temporal lifetime: total elapsed time between first and last consumption of an item.



Item Lifetimes Theoretical Analysis

For simple "copying" model with recency only, we can analyze conditions in which an item lives forever:

Theorem:

Let lpha be probability of novel item $ext{If } \sum_{i=1}^\infty w_i < 1/lpha ext{ then } \Pr[ext{lifetime}(x) < \infty] o 1$



Before items are abandoned, the gap between consumptions of that item grows in both "index" and "real" time.





Consider a simplified choice model with uniform time and item quality scores.

Theorem: Suppose that the weights *w* are monotonically decreasing. Then:

- $1. \quad E[j^{th}\mathrm{gap}] < E[(j-1)^{st}\mathrm{gap}]$
- 2. $E[j^{th}gap|last occurrence] > E[j^{th}gap]$
- $3. \hspace{0.2cm} orall j > J_0: E[j^{th} \mathrm{gap} | j^{th} ext{ is last}] > E[j-1^{st} \mathrm{gap} | j^{th} ext{ is last}]$

Parsimonious model

- Recency weights can be compressed
- Good fit: power law with exponential cutoff:



 $\Pr[x] \propto (x+c)^{-a} e^{-bx}$

Parsimonious model

Dataset	Recency@50	PLECO
BrightKite	0.654	0.926
GPLUS	0.710	0.987
MAPCLICKS	0.668	0.921
WIKICLICKS	0.971	0.999
YouTube	0.917	0.997

Recency model can be expressed using just three parameters!

Satiation



No evidence of satiation in online user behavior

Additivity assumption

Very small deviations from additive behavior

Mildly superadditive as popular items chosen

 $\frac{w_i+w_j}{w(i,j)}$

Getting addicted: superadditive?

Getting bored: subadditive?



Tipping behavior

In the recency model, tipping occurs if after a certain time, only one item is repeatedly consumed

Assume weights are decreasing: $w(p) \ge w(p+1)$

Claim. If sum of weights is finite, then tipping occurs with constant probability

Claim. If the sum of weights is infinite, then tipping does not occur

Conclusions

We studied a number of algorithmic problems related to discrete choice We believe this class of problems is theoretically important and relevant in practice Some open questions

Can one reconstruct, with poly(n) max-sample queries, the winning probabilities of all slates with $o(1) \ell_1$ -error?

What is the relative power of the max-sample / max-dist oracles?

How well can one approximate general mixtures of MNLs with the two oracles?

Identifiability of non-uniform 2-MNLs, k-MNLs